Lecture 3 Probability - Part 1

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Intro

Bayesian vs Frequentist Interpretations

Probability Theory Review

- Foundations of Probability
- Random Variables
- Discrete Random Variables
- Important Rules of Probability
- Independence and Conditional Independence
- Continuous Random Variables

- Binomial and Bernoulli Distributions
- Multinomial and Multinoulli Distributions
- Poisson Distribution
- Empirical Distribution

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Probability: Bayesian vs Frequentist Interpretations



- what is probability?
- there are actually at least two different interpretations of probability
 - **1** frequentist: probabilities represent long run frequencies of events (trials)
 - Bayesian: probability is used to quantify our uncertainty about something (information rather than repeated trials)
- coin toss event:
 - Irequentist: if we flip the coin many times, we expect it to land heads about half the time
 - Bayesian: we believe the coin is equally likely to land heads or tails on the next toss
- advantage of the Bayesian interpretation: it can be used to model our uncertainty about events that do not have long term frequencies; frequentist needs repetition
- the basic rules of probability theory are the same, no matter which interpretation is adopted

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In order to define a **probability space** we need 3 components $\{\Omega, \mathcal{F}, P\}$:

- sample space Ω: the set of all the outcomes of a random experiment. Here, each outcome (realization) ω ∈ Ω can be thought of as a *complete description of the state of the real world* at the end of the experiment
- event space *F*: a set whose elements *A* ∈ *F* (called events) are subsets of Ω (i.e., *A* ⊆ Ω is a collection of possible outcomes of an experiment)
 F should satisfy 3 properties (*σ*-algebra of events):

$$\begin{array}{l} \textcircled{0} & \emptyset \in \mathcal{F} \\ \textcircled{0} & A \in \mathcal{F} \Rightarrow \overline{A} = \Omega \setminus A \in \mathcal{F} \\ \textcircled{0} & A_1, A_2, \dots \in \mathcal{F} \Rightarrow \cup_i A_i \in \mathcal{F} \end{array} (closure under complementation) \\ \begin{array}{l} \textcircled{0} & (closure under complementation) \\ \hline{0} & (closure under countable union) \end{array}$$

probability measure P: a function P : F → R that satisfies the following 3 axioms of probability



experiment: tossing a six-sided dice

- sample space $\Omega = \{1,2,3,4,5,6\}$
- trivial event space

•
$$\mathcal{F} = \{\emptyset, \Omega\}$$

• unique probability measure satisfying the requirements is given by $P(\emptyset)=0, P(\Omega)=1$

- power set event space
 - $\mathcal{F} = 2^{\Omega}$ (i.e., the set of all subsets of Ω)
 - a possible probability measure P(i) = 1/6 for $i \in \{1, 2, 3, 4, 5, 6\} = \Omega$

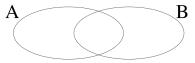
question: do the above sample space outcomes completely describe the state of a dice-tossing experiment?

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Probability Measure Properties

some important properties on events (can be inferred from axioms)

- $A \subseteq B \Rightarrow P(A) \leq P(B)$
- $P(A \cap B) \leq min(P(A), P(B))$
- union bound: $P(A \cup B) \leq P(A) + P(B)$
- complement rule: $P(\overline{A}) = P(\Omega \setminus A) = 1 P(A)$
- impossible event: $P(\emptyset) = 0$
- law of total probability: if $A_1, ..., A_k$ are a set of disjoint events such that $\bigcup_{i=1}^N A_i = \Omega$ then $\sum_{i=1}^N P(A_i) = 1$



• general addition rule¹: $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

 $^1\text{events}$ can be represented by using Venn diagrams $$\$ < ${}_{\odot}$ > <

- let B be an event with non-zero probability, i.e. p(B) > 0
- the conditional probability of any event A given B is defined as

$$P(A|B) = \frac{p(A \cap B)}{p(B)}$$

- in other words, P(A|B) is the probability measure of the event A after observing the occurrence of event B
- two events are called independent iff

 $P(A \cap B) = P(A)P(B)$ (or equivalently P(A|B) = P(A))

• therefore, **independence** is equivalent to saying that observing *B* does not have any effect on the probability of *A*

Conditional Probability

a frequentist intuition of conditional probability

- N is total number of experiment trials
- for an event *E*, let's define $P(E) \triangleq \frac{N_E}{N}$ where N_E is the number of trials where *E* is verified

hence for events A and B (considering the limit $N \to \infty$)

- $P(A) = \frac{N_A}{N}$ where N_A is the number of trials where A is verified
- $P(B) = \frac{N_B}{N}$ where N_B is the number of trials where B is verified
- $P(A \cap B) = \frac{N_{A \wedge B}}{N}$ where $N_{A \wedge B}$ is the number of trials where both A and B are verified

let's consider only the trials where B is verified, hence

• $P(A|B) = \frac{N_{A \wedge B}}{N_B}$ ($N_B > 0$ now acts as N)

• dividing by N, one obtains $P(A|B) = \frac{N_{A \wedge B}/N}{N_B/N} = \frac{P(A \cap B)}{P(B)}$

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intuition: a random variable represents an interesting "aspect" of the outcomes $\omega\in\Omega$

more formally:

- a random variable X is a function $X : \Omega \to \mathbb{R}$
- a random variable is denoted by using upper case letters X(ω) or more simply X (here X is a function)
- the particular values (instances) of a random variable may take on are denoted by using lower case letters x (here x ∈ ℝ)

types of random variables:

- discrete random variable: function $X(\omega)$ can only take values in a finite set $\mathcal{X} = \{x_1, x_2, ..., x_m\}$ or countably infinite set (e.g. $\mathcal{X} = \mathbb{N}$)
- continuous random variable: function $X(\omega)$ can take continuous values in \mathbb{R}

a random variable is a measurable function

- since X(ω) takes values in ℝ, let's try to define an "event space" on ℝ: in general we would like to observe if X(ω) ∈ B for some subset B ⊂ ℝ
- as "event space" on \mathbb{R} , we can consider \mathcal{B} the Borel σ -algebra on the real line², which is generated by the set of half-lines $\{(-\infty, a] : a \in (-\infty, \infty)\}$ by repeatedly applying union, intersection and complement operations
- an element $B \subset \mathbb{R}$ of the Borel σ -algebra \mathcal{B} is called a **Borel set**
- $\bullet\,$ the set of all open/closed subintervals in ${\mathbb R}$ are contained in ${\mathcal B}$
- for instance, $(a, b) \in \mathcal{B}$ and $[a, b] \in \mathcal{B}$
- a random variable is a **measurable function** $X : \Omega \to \mathbb{R}$, i.e.

$$X^{-1}(B) = \{\omega \in \Omega : X(\omega) \in B\} \in \mathcal{F}$$
 for each $B \in \mathcal{B}$

i.e., if we consider an "event" $B \in \mathcal{B}$ this can be represented by a proper event $F_B \in \mathcal{F}$ where we can apply the probability measure P

²here we should use the notation $\mathcal{B}(\mathbb{R})$, for simplicity we drop \mathbb{R}

we have defined the probability measure P on F, i.e. P : F → R
how to define the probability measure P_X w.r.t. X?

$$P_X(B) \triangleq P(X^{-1}(B)) = P(\{\omega \in \Omega : X(\omega) \in B\})$$

which is well-defined given that $X^{-1}(B) \in \mathcal{F}$

• at this point, we have an **induced probability space** $\{\Omega_X, \mathcal{F}_X, P_X\} \triangleq \{\mathbb{R}, \mathcal{B}, P_X\}$ and we can equivalently reason on it

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Discrete Random Variables

discrete Random Variable (RV)

- X(ω) can only take values in a finite set X = {x₁, x₂, ..., x_m} or in a countably infinite set
- how to define the probability measure P_X w.r.t. X?

$$P_X(X = x_k) \triangleq P(\{\omega : X(\omega) = x_k\})$$

- in this case P_X returns measure one to a countable set of reals
- a simpler way to represent the probability measure is to directly specify the probability of each value the discrete RV can assume
- in particular, a **Probability Mass Function** (PMF) is a function $p_X : \mathbb{R} \to \mathbb{R}$ such that

$$p_X(X=x) \triangleq P_X(X=x)$$

• it's very common to drop the subscript X and denote the PMF with $p(X) = p_X(X = x)$

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considering two discrete RV X and Y at the same timesum rule

$$p(X) = \sum_{Y} p(X, Y)$$
 (marginalization)

product rule

$$p(X,Y) = p(X|Y)p(Y)$$

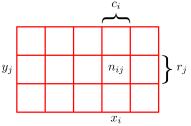
• chain rule:

 $p(X_{1:D}) = p(X_1)p(X_2|X_1)p(X_3|X_2,X_1)...p(X_D|X_{1:D-1})$

where 1: D denotes the set $\{1, 2, ..., D\}$ (Matlab-like notation)

Important Rules of Probability





- N number of trials
- n_{ij} number of trials in which $X = x_i$ and $Y = y_j$
- c_i number of trials in which $X = x_i$, one has $c_i = \sum_j n_{ij}$
- r_j number of trials in which $Y = y_j$, one has $r_j = \sum_i n_{ij}$

• $p(X = x_i, Y = y_j) = \frac{n_{ij}}{N}$ (considering the limit $N \to \infty$)

hence:

•
$$p(X = x_i) = \frac{c_i}{N} = \sum_j \frac{n_{ij}}{N} = \sum_j p(X = x_i, Y = y_j)$$

combining the definition of condition probability with the product and sum rules:

a $p(X|Y) = \frac{p(X,Y)}{P(Y)}$ (conditional prob. def.) **b** p(X,Y) = p(Y|X)p(X) (product rule) **c** $p(Y) = \sum_{X} p(X,Y) = \sum_{X} p(Y|X)p(X)$ (sum rule + product rule)

one obtains the Bayes' Theorem

$$p(X|Y) = \frac{p(Y|X)p(X)}{\sum_{X} p(Y|X)p(X)}$$

N.B.: we could write $p(X|Y) \propto p(Y|X)p(X)$; the denominator $p(Y) = \sum_{X} p(Y|X)p(X)$ can be considered as a normalization constant events:

- C = breast cancer present, \overline{C} = no cancer
- M = positive mammogram test, $\overline{M} = \text{negative mammogram test}$

probabilities:

- p(C) = 0.4% (hence $p(\overline{C}) = 1 p(C) = 99.6\%$)
- if there is cancer, the probability of a pos mammogram is p(M|C) = 80%
- if there is no cancer, we still have $p(M|\overline{C}) = 10\%$

false conclusion: positive mammogram \Rightarrow the person is 80% likely to have cancer question: what is the conditional probability p(C|M)?

$$p(C|M) = \frac{p(M|C)p(C)}{p(M)} = \frac{p(M|C)p(C)}{p(M|C)p(C) + p(M|\overline{C})p(\overline{C})}$$
$$= \frac{0.8 \times 0.004}{0.8 \times 0.004 + 0.1 \times 0.996} = 0.031$$

true conclusion: positive mammogram \Rightarrow the person is about 3% likely to have cancer

Image: A matrix and a matrix

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Independence and Conditional Independence

considering two RV X and Y at the same time

• X and Y are unconditionally independent

$$X \perp Y \iff p(X,Y) = p(X)p(Y)$$

in this case p(X|Y) = p(X) and p(Y|X) = p(Y)

• X₁, X₂, ..., X_D are mutually independent if

$$p(X_1, X_2, ..., X_D) = p(X_1)p(X_2)...p(X_D)$$

• X and Y are conditionally independent

$$X \perp Y | Z \iff p(X, Y | Z) = p(X | Z) p(Y | Z)$$

in this case p(X|Y,Z) = p(X|Z)

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continuous random variable

- $X(\omega)$ can take any value on $\mathbb R$
- how to define the probability measure P_X w.r.t. X?

$$P_X(X \in B) \triangleq P(X^{-1}(B))$$
 (with $B \in \mathcal{B}$)

• in this case P_X gives zero measure to every singleton set, and hence to every countable set³

 3 unless we consider some particular/degenerate cases \prec = > \prec = > \rightarrow

given a continuous RV X

- Cumulative Distribution Function (CDF): $F(x) \triangleq P_X(X \le x)$
 - $0 \leq F(x) \leq 1$
 - the CDF is a monotonically non-decreasing $F(x) \leq F(x + \Delta x)$ with $\Delta x > 0$

•
$$F(-\infty) = 0, F(\infty) = 1$$

•
$$P_X(a < X \le b) = F(b) - F(a)$$

• **Probability Density Function** (PDF): $p(x) \triangleq \frac{dr}{dx}$

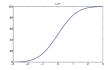
• we assume F is continuous and the derivative exists

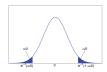
•
$$F(x) = P_X(X \le x) = \int_{-\infty}^{x} p(\xi) d\xi$$

•
$$P_X(x < X \le x + dx) \approx p(x)dx$$

•
$$P_X(a < X \le b) = \int_a^b p(x) dx$$

p(x) acts as a density in the above computations





reconsider

- $P_X(a < X \le b) = \int_a^b p(x) dx$
- $P_X(x < X \le x + dx) \approx p(x)dx$

• the first implies
$$\int_{-\infty}^{\infty} p(x) dx = 1$$

(consider
$$(a, b) = (-\infty, \infty)$$
)

- the second implies $p(x) \ge 0$ for all $x \in \mathbb{R}$
- it is possible that p(x) > 1, for instance, consider the **uniform distribution** with PDF

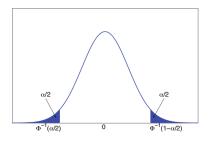
$$\mathsf{Unif}(x|a,b) = \frac{1}{b-a}\mathbb{I}(a \le x \le b)$$

if a = 0 and b = 1/2 then p(x) = 2 in [a, b]

assume F is continuous (this was required for defining p(x))
we have that P_X(X = x) = 0 (zero probability on a singleton set)
in fact for ε ≥ 0: P_X(X = x) ≤ P_X(x - ε < X ≤ x) = F(x) - F(x - ε) = δF(x, ε) and given that F is continuous P_X(X = x) ≤ lim_{ε→0} δF(x, ε) = 0

Quantile

- given that the CDF F is monotonically increasing, let's consider its inverse F^{-1}
- $F^{-1}(\alpha) = x_{\alpha} \iff P_X(X \le x_{\alpha}) = \alpha$
- x_{α} is called the α quantile of F
- $F^{-1}(0.5)$ is the median
- $F^{-1}(0.25)$ and $F^{-1}(0.75)$ are the lower and upper quartiles
- for symmetric PDFs (e.g. N(0,1)) we have F⁻¹(1 − α/2) = −F⁻¹(α/2) and the central interval (F⁻¹(α/2), F⁻¹(1 − α/2)) contains 1 − α of the mass probability



• mean or expected value μ

for a discrete RV: $\mu = \mathbb{E}[X] \triangleq \sum_{x \in \chi} x \ p(x)$

for a continuos RV: $\mu = \mathbb{E}[X] \triangleq \int_{x \in \chi} x \ p(x) \ dx$ (defined if $\int_{x \in \chi} |x| \ p(x) \ dx < \infty$)

• variance
$$\sigma^2 = \operatorname{var}[X] \triangleq \mathbb{E}[(X - \mu)^2]$$

$$\operatorname{var}[X] = \mathbb{E}[(X - \mu)^2] = \int_{x \in \chi} (x - \mu)^2 p(x) dx =$$
$$= \int_{x \in \chi} x^2 p(x) dx - 2\mu \int_{x \in \chi} x p(x) dx + \mu^2 \int_{x \in \chi} p(x) dx = \mathbb{E}[X^2] - \mu^2$$

(this can be also obtained for discrete RV)

• standard deviation $\sigma = \operatorname{std}[X] = \sqrt{\operatorname{var}[X]}$

• *n*-th moment

for a discrete RV: $\mathbb{E}[X^n] \triangleq \sum_{x \in \chi} x^n p(x)$

for a continuos RV: $\mathbb{E}[X^n] \triangleq \int_{x \in Y} x^n p(x) dx$ (defined if $\int_{x \in Y} |x|^n p(x) dx < \infty$)

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Binomial Distribution

- we toss a **coin** *n* times
- X is a discrete RV with $x \in \{0, 1, ..., n\}$, the occurred number of heads
- θ is the probability of heads
- $X \sim Bin(n, \theta)$ i.e., X has a **binomial distribution** with PMF

$$\mathsf{Bin}(k|n,\theta) \triangleq \binom{n}{k} \theta^k (1-\theta)^{n-k} \qquad (= P_X(X=k))$$

where we use the binomial coefficient

$$\binom{n}{k} \triangleq \frac{n!}{(n-k)!k!}$$

• mean = $n\theta$

- var = $n\theta(1-\theta)$
- N.B.: recall that $(a + b)^n = \sum_{k=0}^n {n \choose k} a^k b^{n-k}$

- we toss a coin only one time
- X is a discrete RV with $x \in \{0,1\}$ where 1 = head, 0 = tail
- θ is the probability of heads
- $X \sim Ber(\theta)$ i.e., X has a **Bernoulli distribution** with PMF

$$\mathsf{Ber}(x| heta) \triangleq heta^{\mathbb{I}(x=1)}(1- heta)^{\mathbb{I}(x=0)} \qquad \qquad (=P_X(X=x))$$

that is

$$\mathsf{Ber}(x|\theta) = \begin{cases} \theta & \text{if } x = 1\\ 1 - \theta & \text{if } x = 0 \end{cases}$$

• mean = θ

• var = $\theta(1-\theta)$

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Multinomial Distribution

- we toss a K-sided dice n times
- the possible outcome is x = (x₁, x₂, ..., x_K) where x_j ∈ {0, 1, ..., n} is the number of times side j occurred
- $n = \sum_{j=1}^{K} x_j$
- θ_j is the probability of having side j

•
$$\sum_{j=1}^{K} \theta_j = 1$$

• $X \sim Mu(n, \theta)$ i.e., X has a multinomial distribution with PMF

$$\mathsf{Mu}(\mathbf{x}|n,\boldsymbol{\theta}) \triangleq \binom{n}{x_1 \dots x_K} \prod_{j=1}^K \theta_j^{x_j}$$

where we use the multinumial coefficient

$$\binom{n}{x_1 \dots x_K} \triangleq \frac{n!}{x_1! x_2! \dots x_K!}$$

which is the num of ways to divide a set of size n into subsets of size $x_1, x_2, ..., x_K$

- we toss the dice only one time
- the possible outcome is $\mathbf{x} = (\mathbb{I}(x_1 = 1), \mathbb{I}(x_2 = 1), ..., \mathbb{I}(x_K = 1))$ where $x_j \in \{0, 1\}$ represents if side j occurred or not (dummy enconding or one-hot encoding)
- θ_j is the probability of having side *j*, i.e., $p(x_j = 1|\theta) = \theta_j$
- $X \sim Cat(\theta)$ i.e., X has the categorical distribution (or multinoulli)

$$\mathsf{Cat}(\mathsf{x}|\boldsymbol{ heta}) = \mathsf{Mu}(\mathsf{x}|1, \boldsymbol{ heta}) riangleq \prod_{j=1}^{K} heta_{j}^{\mathsf{x}_{j}}$$

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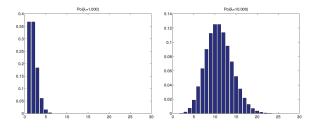
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Poisson Distribution

- X is a discrete RV with $x \in \{0, 1, 2, ...\}$ (support on \mathbb{N}^+)
- $X \sim \text{Poi}(\lambda)$ i.e., X has a **Poisson distribution** with PMF

$$\operatorname{Poi}(x|\lambda) \triangleq e^{-\lambda} \frac{\lambda^{x}}{x!}$$

- recall that $e^{\lambda} = \sum_{x=0}^{\infty} \frac{\lambda^x}{x!}$
- this distribution is used as a model for counts of rare events (e.g. accidents, failures, etc)



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Empirical Distribution

- given a dataset $\mathcal{D} = \{x_1, x_2, ..., x_N\}$
- the empirical distribution is defined as

$$p(x) = \sum_{i=1}^{N} w_i \delta_{x_i}(x)$$

•
$$0 \le w_i \le 1$$
 are the weights

•
$$\sum_{i=1}^{N} w_i = 1$$

•
$$\delta_{x_i}(x) = \mathbb{I}(x = x_i)$$

• this can be view as an **histogram** with "spikes" at $x_i \in D$ and 0-probability out D

- Kevin Murphy's book
- A. Maleki and T. Do "*Review of Probability Theory*", Stanford University
- G. Chandalia "A gentle introduction to Measure Theory"